

Is the observed NAO variability during the instrumental record unusual?

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Received 12 January 2008; revised 24 March 2008; accepted 29 April 2008; published 5 June 2008.

[1] Observed multidecadal variability (30 yr running means, trends, and moving standard deviations) of the North Atlantic Oscillation (NAO) during the instrumental record is compared to that simulated by two different coupled general circulation models in extended-range control experiments. Simulated NAO exhibits strong low frequency fluctuations, even on multi-centennial time scale. Observed multi-decadal NAO variations agree well with the model variability. Trend probability distribution functions, observed and simulated, were not found to be different with statistical significance. Thus, multi-decadal NAO changes similar to those observed during the instrumental record, including the recent increase in 1965–1995, may be internally generated within the coupled atmosphere-ocean system without considering external forcing. **Citation:** Semenov, V. A., M. Latif, J. H. Jungclauss, and W. Park (2008), Is the observed NAO variability during the instrumental record unusual?, *Geophys. Res. Lett.*, 35, L11701, doi:10.1029/2008GL033273.

1. Introduction

[2] The North Atlantic Oscillation (NAO), the dominant mode of large-scale wintertime atmospheric variability in the North Atlantic/European sector [Walker, 1924; van Loon and Rogers, 1978], has received a lot of attention due to its strong impact on climate variability in the northern Extra-Tropics. In particular, European surface air temperatures (SAT) and precipitation are strongly related to the NAO [Hurrell, 1995]. The NAO experienced significant decadal and multi-decadal changes during the 20th century that were reflected in different components of the climate system (see, e.g., Hurrell *et al.* [2003] for a review).

[3] The strongest NAO trend during the instrumental period was observed during 1965–1995 (Figure 1a). In the presence of strong global climate change during the last half of the 20th century, it is tempting to explain the NAO trend by forcing external to the atmosphere. A number of studies linked the NAO trend to North Atlantic (NA) and/or tropical sea surface temperatures (SSTs), which exhibited similar contemporaneous changes. Atmosphere model simulations show SSTs (local or remote) may have driven the decadal to multi-decadal NAO changes (see, e.g., Hoerling *et al.* [2004] for a review). However, some models contradict each other on the relative roles of local and remote forcing, while others do not reproduce the NAO trend at all

[e.g., Schneider *et al.*, 2003]. Furthermore, most model studies concentrate only on the second half of the 20th century (covered by reanalysis) and simulated NAO trends are too weak. Thus, the results should be treated with caution [Bretherton and Battisti, 2000].

[4] In Figure 1, the NAO index is shown together with an Indian Ocean SST index (15°S–15°N, 40°E–100°E), as by Hoerling *et al.* [2004], and an index of the NA SST tripole regression pattern associated with the NAO [e.g., Rodwell *et al.*, 1999], calculated by regressing detrended SST against the NAO index for 1950–2004. The NA SST tripole index was computed as the scalar product between a regression pattern and SST anomalies in the region 10°N–70°N, 80°W–0°W. The positive NAO trend during 1965–1995 was accompanied by a warming of the Indian Ocean. The downward NAO trend during 1910–1940, however, was also accompanied by a *positive SST trend* in the Indian Ocean comparable in magnitude to the recent warming. Thus, no robust link seems to exist when considering the whole 20th century. Furthermore, the NA SST tripole regression index for 1950–2004 is representative only for the time after 1960. For the earlier period no robust connection is seen either. This is a simple illustration of some questions that remain after attributing the recent NAO changes to SST forcing.

[5] A shift to the positive NAO phase is generally simulated in (anthropogenic) climate change experiments, although the changes are rather small amounting to only 10–15% of the standard deviation [Kuzmina *et al.*, 2005; Stephenson *et al.*, 2006]. Corresponding global warming in these simulations, however, is much stronger than the 20th century temperature rise. Transient climate change integrations assuming the moderate Intergovernmental Panel on Climate Change (IPCC) scenario B1, for instance, yield a range of global warming of about 1.2°C to 2.4°C (with an average of 1.76°C) at the end of 21st century [IPCC, 2007], which is at least 3 times larger than the global warming observed since 1950s (0.4°C). Given the relatively small change of the NAO in climate change experiments, a potential anthropogenic induced NAO change during the 20th century is unlikely to be detected against the strong background noise.

[6] The instrumental NAO record basically consists of only one realization of the multi-decadal variability (Figure 1). One way to estimate the level of natural variability on multi-decadal time scales is to make use of long simulations with realistic coupled general circulation models (GCMs) run in a “control” setup, i.e. without external forcing. So far, not many simulations exceeding 1000 years were performed. The NAO in millennium control simulations is considered, for instance, by Osborn *et al.* [1999], Collins *et al.* [2001], and Min *et al.* [2005]. Here, the NAO variability is

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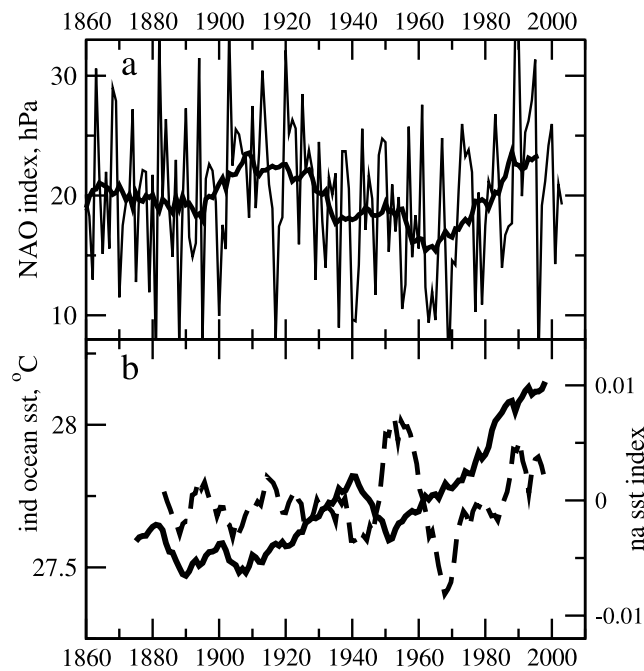


Figure 1. (a) The observed NAO index (DJFM SLP difference [hPa] between Gibraltar and Reykjavik [Jones *et al.*, 1997]. The thick line is the 15-year running means. (b) Indian Ocean SST index (15S–15N, 40E–100E average, 15-year running means, solid line), and index of the NA SST regression pattern (15-year running means, dashed; see text for details).

analyzed in control simulations with two other coupled GCMs, ECHAM5/MPI-OM and Kiel Climate Model (KCM), 1500 years and 3150 years long, respectively. We focus on the multi-decadal changes of the NAO characteristics and address the question of how unusual were the NAO trends during the observational record by applying a test on probability distribution functions.

2. Models and Data

[7] Two coupled atmosphere-ocean GCMs are used in this study. The first is ECHAM5/MPI-OM (hereafter referred to as MPI) [Jungclaus *et al.*, 2006], and the other is the Kiel Climate Model, KCM (W. Park *et al.*, Tropical Pacific climate and its response to global warming, submitted to *Journal of Climate*, 2008). Both models use the atmospheric GCM ECHAM5 [Roeckner *et al.*, 2003] run at a spectral resolution of T31, corresponding to a horizontal resolution of about $3.75^\circ \times 3.75^\circ$, and with 19 vertical levels. The two atmospheric model versions used in the simulations with MPI and KCM differ in the cloud parameterization scheme [Roeckner *et al.*, 2003]. MPI employs the ocean model MPI-OM [Marsland *et al.*, 2003]. This is a primitive equation model (z-level, free surface) on a C-grid with variable horizontal resolution between 20 km in high latitudes and about 350 km in the Tropics, and 40 vertical levels. KCM uses the oceanic GCM NEMO [Madec, 2006]. It is also a z-coordinate model (31 vertical levels) with partial bottom cells and a free surface formulation. The

horizontal resolution is based on a 2° Mercator mesh and is on average 1.3° , with enhanced meridional resolution of 0.5° close to the equator (ORCA2).

[8] Both coupled atmosphere-ocean GCMs do not employ flux adjustments or any other corrections. The MPI and KCM control simulations are of 1700 and 3750 year duration, respectively. To account for spin-up, only the last 1500 years of the MPI and 3150 years of the KCM simulation are analyzed. Some examples of the models' ability to realistically simulate the internal variability are given by Semenov and Latif [2006] and Dommenget *et al.* [2006].

[9] Here we use absolute sea level pressure (SLP) difference between Gibraltar and Reykjavik averaged for December through March (DJFM) as the NAO index. This index, starting in 1824, is the longest time series based on instrumental observations [Jones *et al.*, 1997]. The data for 1824–2003 are available from the Climate Research Unit (CRU, <http://www.cru.uea.ac.uk/cru/data/nao.htm>).

[10] Ability of the models to realistically simulate atmospheric variability in the northern Extra-Tropics was validated using NCEP reanalysis (1949–2006) [Kalnay *et al.*, 1996]. The longer historical (gridded) SLP dataset HadSLP (1850–2004) [Allan and Ansell, 2006] was not used as these data systematically exhibit about 20% less interannual variability than NCEP. Standard deviations of the DJFM SLP for NCEP and the models are shown in the Figure 2. Both models simulate SLP variability patterns consistent with NCEP. Although the simulated variability is overestimated by the models in the North Pacific and underestimated over Asia and the Tropical Atlantic, no statistically significant disagreement (according to an F-test) is found in the North Atlantic Sector. The standard deviation of the NAO index amounts to 6.3 hPa in station data, 6.8 hPa in NCEP, 5.1 hPa in HadSLP, 6.4 hPa in MPI, and 6.7 hPa in KCM.

3. Results

[11] Running 30-year means, trends, and standard deviations (in a sliding 30-year window) of the NAO index as computed from observations and model simulations are analyzed. Only the results of the longer KCM simulation are shown in Figure 3. The MPI simulation provides a very similar picture and will be discussed later in the context of the trend distributions. The observed 30-yr running mean NAO changes are within the variability range simulated in the KCM control experiment (Figure 3a). The simulated NAO index exhibits several realizations that are similar to the observed multi-decadal fluctuations. Even on the centennial time scale, the simulated NAO variability reveals strong changes. We note, for example, a strong secular decrease around the model year 4800, or exceptionally high values around the year 2900.

[12] The moving 30-year trends are depicted in Figure 3b. The strongest observed positive trend is slightly exceeded once during the KCM simulation. Although the strongest observed trend (0.40 hPa/yr) corresponds to the extreme outliers of the simulated trends (exceeding the 99% percentile; see Figure 3b), the simulated trend variability, e.g. during the first 500 years, looks similar to the observed.

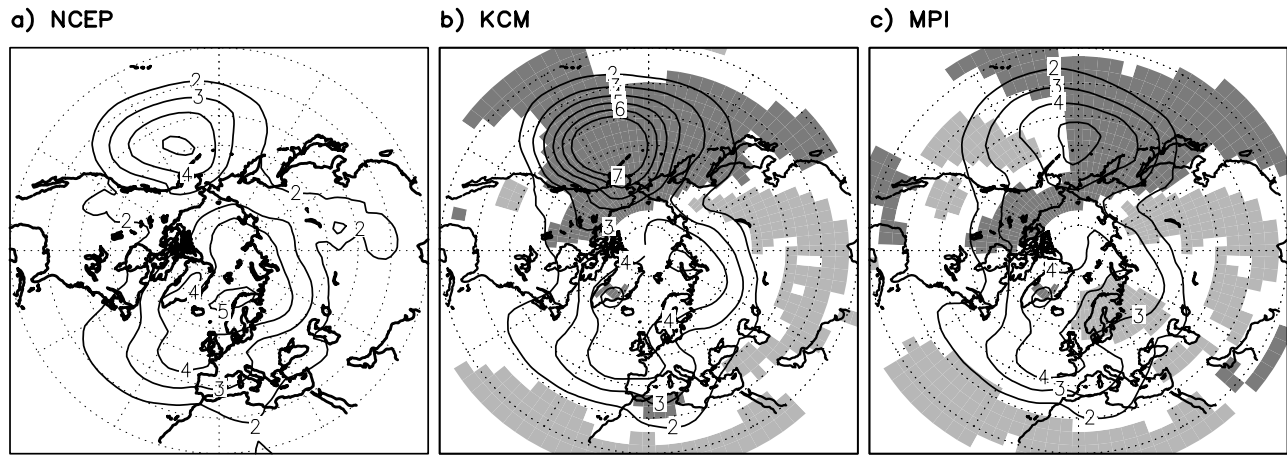


Figure 2. Standard deviation of the DJFM SLP variability [hPa] from (a) NCEP, (b) KCM, and (c) MPI. The shading shows regions where variability ratio model/NCEP is statistically larger (dark shading) or lower (light shading) at the 95% confidence level according to an F-test.

Furthermore, one should also consider the simulated negative trends, which exceed the strongest observed trend in terms of absolute amplitude in several cases.

[13] It was found that the variance of the NAO also increased with the recent positive trend [Feldstein, 2002]. The role of global warming in this observation remains speculative. The moving 30-year standard deviations (STD) of the simulated NAO index ranges from 3.8 hPa to 9.6 hPa encompassing the observed changes and exhibit strong long-term variability (Figure 3c). The time-mean STDs for

the whole periods are, however, very close, with 6.3 hPa and 6.8 hPa for the observed and model index, respectively. A comparison with the smoothed NAO index (Figure 3a) does not reveal any noticeable link to the variance (no significant correlation is found between running means and STD).

[14] An externally forced nature of the observed positive NAO trend was inferred in several studies [e.g., Osborn *et al.*, 1999; Collins *et al.*, 2001; Feldstein, 2002; Min *et al.*, 2005] through comparison with trend distributions obtained

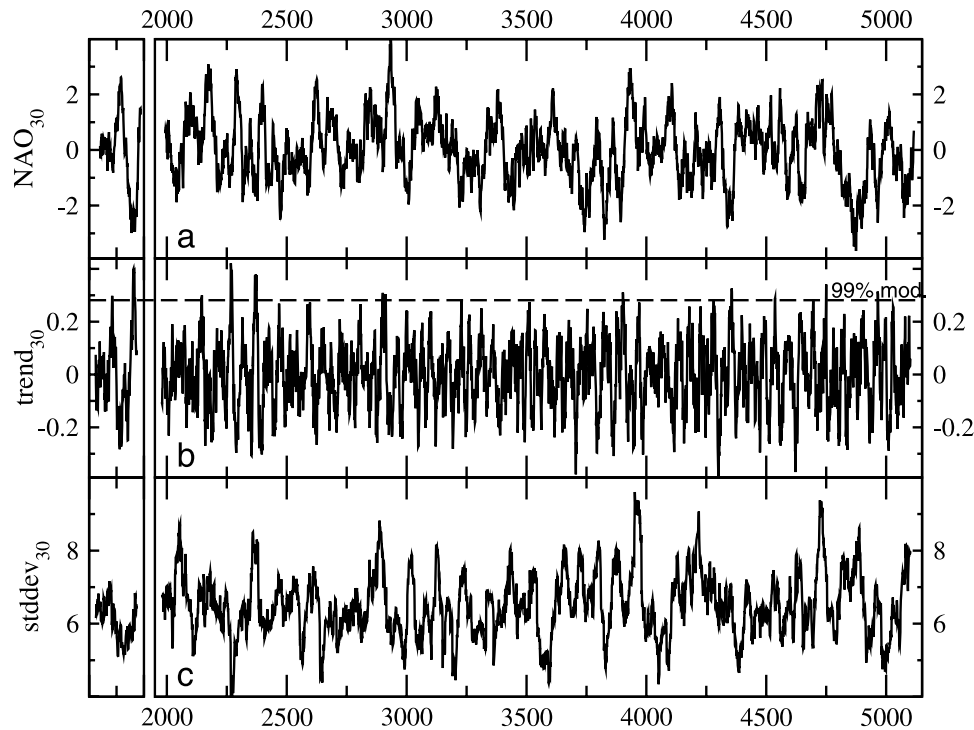


Figure 3. (a) The 30-year running means of the NAO index [hPa] as (left) observed and (right) simulated by KCM, (b) moving 30-year trends [hPa/yr] (dashed line is the 99% percentile of the model trend distribution), and (c) standard deviation of the NAO index [hPa] in a moving 30-year window.

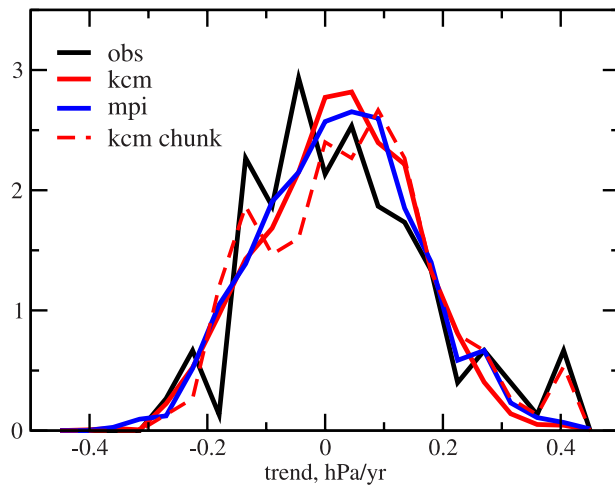


Figure 4. Probability density functions (histograms) of the observed (black), KCM (red solid), and MPI (blue) simulated 30-year moving NAO trends. The dashed red histogram is for a particular 170 year chunk from the KCM simulation (model years 2131–2280).

from control climate model simulations or statistical models. The null hypothesis that the observed trend was an expression of stochastic natural variability was rejected based on a rare occurrence (or absence) of the largest observed trend among simulated values or, for instance, a location of the observed trend beyond the 95% percentile of the model trend distribution. However, models simulate multi-decadal variations in NAO regimes rather similar to the observed changes in the last two centuries (Figure 3). Thus, we shall test not the unusualness of the strongest observed trend in comparison to the simulated trend distribution, but the unusualness of the observed trend distribution for the whole observational period. When looking at the probability density functions (PDFs) of the observed and simulated 30-year trends (Figure 4), it is hard to conclude with high confidence that they may not be drawn from the same distribution. Rather, taking into account the short length of the observational record, the differences may be caused just by under-sampling. Extreme positive outliers look as suspicious as a jagged behavior around zero or a drop at about -0.2 . This is illustrated by the KCM trend distribution for model years 2131–2280 (same length as observed NAO), which also shows an upper tail spike.

[15] In order to objectively test the difference between the observed and model PDFs, a Kolmogorov-Smirnov test (K–S) [e.g., *D’Agostino and Stephens*, 1986] is applied, measuring the maximum difference between two empirical distribution functions. The null hypothesis tested is that observed and model distributions are not different or equivalently that observed and model samples are drawn from the same distribution. The K–S test returns probabilities (p values) of 0.39 and 0.62 for KCM and MPI, respectively (small p values are required to reject the null hypothesis). Thus the null hypothesis in our case may not be rejected. The moving 30-year observed trend distribution, however, contains only six fully independent points. The K–S test was applied also to the distributions of moving, consecutive and overlapping trends of 30 and 20 year

lengths. The results were found to depend on the method used to construct the PDFs. However, the test also does not yield significance levels less than 0.2.

4. Discussion

[16] It is shown that the long-term NAO changes observed during the instrumental record including the recent strongest positive trends may be reproduced in control (without external forcing) simulations with coupled atmosphere-ocean GCMs (Figure 3). Simulated NAO indices exhibit significant changes on multi-decadal and even multi-centennial time scales. This strongly suggests that longer millennium time scale control simulations (at least of several thousand years) are needed in order to infer the range of natural NAO variability, and longer and reliable NAO reconstructions from paleo archives are very important. These conclusions are crucially dependent on the models’ ability to realistically simulate real climate variability. Here, we base our confidence in the model results on a reasonable agreement of the simulated DJFM SLP variability in comparison to the NCEP reanalysis (Figure 2). The largest observed trend exceeds the 99% percentile of the simulated trend distribution for both control simulations considered here, confirming a very low probability of such a trend. In this respect, our results do not contradict other studies. We choose, however, a different perspective to interpret the unusualness of the long-term observed NAO changes. Based on a goodness-of-fit test applied to the observed and simulated trend distributions, we argue that the null hypothesis that the observed sample of climatic NAO trends and simulated trends (by the KCM and MPI models) are drawn from the same distribution may not be rejected. This is in agreement with conclusions by *Wunsch* [1999] and *Jung et al.* [2003], who tested the recent NAO changes against the “white noise” null hypothesis. The result, however, does not disprove the hypothesis of external forcing, which is not considered here.

[17] **Acknowledgments.** The authors would like to thank two anonymous reviewers and Dr. Noel Keenlyside for valuable comments. This work was supported by the Sonderforschungsbereich 460 of the German Science Foundation (DFG), the European Union’s DYNAMITE project, and the Russian Foundation for Basic Research. The model runs were performed at the German Climate Computing Center (DKRZ) and the Computing Center at Kiel University.

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